

Return your written solutions either in person or by email  
to veska.kaarnioja@fu-berlin.de by Monday 20 June, 2022, 12:15

1. Suppose that our quantity of interest is a function  $f : [0, 1] \rightarrow \mathbb{R}$ ,  $f(0) = 0$ , and we have *a priori* knowledge that the function  $f$  may have a large jump at few places in the interval, but the locations of these places are unknown to us. One way to construct a prior in this case is to consider the finite difference approximation of the derivative of  $f$  and assume that it follows an impulse noise probability distribution.

Let  $t_j = j/d$ ,  $j \in \{0, \dots, d\}$ , be a discretization of the computational domain and let  $x_j = f(t_j)$ . Let us consider the problem of drawing a sample from the  $d$ -dimensional prior density

$$\pi(x) = \left(\frac{\alpha}{\pi}\right)^d \prod_{j=1}^d \frac{1}{1 + \alpha^2(x_j - x_{j-1})^2}, \quad \alpha > 0.$$

Define  $u = (u_1, \dots, u_d)^T$  by

$$u_j = x_j - x_{j-1}, \quad j = 1, \dots, d. \quad (1)$$

The new random variable  $u$  is distributed according to the  $d$ -dimensional Cauchy distribution (*without positivity constraint!*)

$$\pi(u) = \left(\frac{\alpha}{\pi}\right)^d \prod_{j=1}^d \frac{1}{1 + \alpha^2 u_j^2}. \quad (2)$$

The components  $u_j$  are independent of each other and can be drawn from a univariate Cauchy distribution using inverse transform sampling. Note that the inverse cumulative distribution function corresponding to a univariate Cauchy density (*without positivity constraint!*) is given by

$$\Phi^{-1}(t) = \frac{1}{\alpha} \tan(\pi(t - \frac{1}{2})), \quad t \in (0, 1).$$

Implement the following algorithm in MATLAB:

- Fix  $d = 1200$ , set  $\alpha = 1$ , and define  $\mathbf{t} = (1:d)'/d$ ;
- Use componentwise inverse transform sampling to draw a realization from the  $d$ -dimensional Cauchy distribution (2).
- Note that (1) implies that we can transform the realization back into the original coordinates  $x$  via

$$x_1 = u_1, \quad x_2 = u_1 + u_2, \quad x_3 = u_1 + u_2 + u_3, \dots, \quad x_d = u_1 + u_2 + u_3 + \dots + u_d.$$

In MATLAB, this can be achieved with the command `x = cumsum(u)`;

- Finally, visualize the sample you obtained using the `plot` command. Plot the realization as a function over the original spatial mesh  $\mathbf{t}$ .

2. Let  $y \in \mathbb{R}^2$  and  $x \in \mathbb{R}$  and

$$y = \begin{pmatrix} 2 \\ 1 \end{pmatrix} x + \eta, \quad \eta \sim \mathcal{N}(0, \gamma^2 I_2),$$

where  $I_2 \in \mathbb{R}^{2 \times 2}$  is an identity matrix. Suppose the prior distribution is given by  $x \sim \mathcal{N}(0, 2)$ . What is the posterior distribution if we observe  $\hat{y} = (1 \ 2)^\top$ ? What is the posterior variance? What happens to posterior distribution and variance under decreasing noise ( $\gamma \downarrow 0$ )?

3. Let  $A \in \mathbb{R}^{k \times d}$ ,  $x \in \mathbb{R}^d$ ,  $y, \eta \in \mathbb{R}^k$ , and consider the linear measurement model with additive noise:

$$y = Ax + \eta.$$

During the lecture, we proved that if  $x$  is endowed with a Gaussian prior distribution  $\mathcal{N}(x_0, \Gamma_{\text{pr}})$ , the noise  $\eta$  is assumed to have the Gaussian distribution  $\mathcal{N}(\eta_0, \Gamma_{\text{n}})$ , and  $x$  and  $\eta$  are mutually independent, then the posterior distribution is Gaussian with posterior covariance

$$\Gamma_{\text{post}} = (\Gamma_{\text{pr}}^{-1} + A^\top \Gamma_{\text{n}}^{-1} A)^{-1} \quad (3)$$

and posterior mean

$$\mu_{\text{post}} = \Gamma_{\text{post}} (A^\top \Gamma_{\text{n}}^{-1} (y - \eta_0) + \Gamma_{\text{pr}}^{-1} x_0). \quad (4)$$

Prove that

$$\Gamma_{\text{post}} = \Gamma_{\text{pr}} - \Gamma_{\text{pr}} A^\top (A \Gamma_{\text{pr}} A^\top + \Gamma_{\text{n}})^{-1} A \Gamma_{\text{pr}} \quad (5)$$

and

$$\mu_{\text{post}} = x_0 + \Gamma_{\text{pr}} A^\top (A \Gamma_{\text{pr}} A^\top + \Gamma_{\text{n}})^{-1} (y - Ax_0 - \eta_0). \quad (6)$$

*Hint:* Use the Sherman–Morrison–Woodbury formula: for *any* conformable matrices  $A, B, C$ , and  $D$  such that  $A$  and  $C$  are invertible (square) matrices, it holds that

$$(A + BCD)^{-1} = A^{-1} - A^{-1} B (C^{-1} + DA^{-1} B)^{-1} DA^{-1},$$

if  $A + BCD$  is invertible (or, equivalently, if  $C^{-1} + DA^{-1} B$  is invertible).

Begin by applying the Sherman–Morrison–Woodbury formula on (3); this should yield the formula (5). The formula (6) can then be proved by plugging the formula (5) into (4) and simplifying the resulting expression.